Hedging with CO_2 allowances: the ECX market

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- The trading scheme for carbon dioxide (CO₂) emission allowances is one of the major steps towards reducing the environmental burden
- Ø Allowances: new asset with derivatives
- Solution CO2 spot and futures contracts for ECX: lack of hedging analysis
- Hedging is analyzed for several commodities markets
- What about hedging using CO₂ contracts? How it behaves?
- Daily hedging with futures: MV hedge ratios estimated conditionally using multivariate GARCH models and unconditionally by OLS and naïve strategies
- Outility gains are considered in order to take into account risk-return considerations

- Calculate for the first time hedge ratios for the CO₂ allowances market.
- Extend the data span considered by previous authors that mostly covered the Phase I period (2005-2007).
- Use both static and dynamic hedging strategies which allows us to compare different specifications.
- Help to identify the internal dynamics of widely traded CO₂ emission allowances, essential in pricing of the contracts, while the implications of the study are expected to be functional for risk managers, individual investors and hedgers dealing with the carbon allowances trading markets.

- Methodology
- Data and Summary Statistics
- Empirical Results
- Conclusions

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$$h_t = rac{Cov_t(s_{t+1}, f_{t+1})}{Var_t(f_{t+1})}$$

 Effectiveness of the hedging strategy measured by: the degree of hedging effectiveness (EH) - variance reduction - and utility maximization - from a utility gains standpoint

• Variance Reduction:

$$EH = \frac{Var(\Delta S_t) - Var(\Delta h_t)}{Var(\Delta S_t)} = 1 - \left[\frac{Variance_{hedgedPortfolio}}{Variance_{unhedgedPortfolio}}\right]$$

• Expected Utility gains: $\begin{aligned} & \underset{h}{Max} U\left[E(r_{p,t}), \sigma_{p,t}; \eta\left(r_{p,t}\right)\right] = \underset{h}{Max} E\left(r_{p,t}\right) - 0, 5\eta\left(r_{p,t}\right) \sigma_{p,t}^{2} ; \\ & \underset{h}{E}[U(r_{p,t})|\psi_{t-1}] = E[r_{p,t}|\psi_{t-1}] - \lambda Var[r_{p,t}|\psi_{t-1}] \end{aligned}$

- When a hedge where the futures position have the same size but the opposite sign than the position held in the spot market is considered, we have what is called a naïve hedge ratio (h_t = 1, ∀t).
- Empirically, the one period hedge ratio is estimated by the slope from the following ordinary least squared (OLS) regression equation:

$$s_{t+1} = \alpha + h^* f_{t+1} + \varepsilon_t$$

where ε_t is the error term from OLS estimation, s_{t+1} and f_{t+1} are the changes in the spot and futures prices, respectively, between time t and t+1, and h^* is the minimum hedge ratio.

Methodology: Time varying hedge ratio

- Multivariate GARCH models
- 1) BEKK

$$\begin{bmatrix} \sigma_{ss,t}^2 & \sigma_{sf,t}^2 \\ \sigma_{fs,t}^2 & \sigma_{ff,t}^2 \end{bmatrix} = \begin{bmatrix} c_{ss} & c_{sf} \\ 0 & c_{ff} \end{bmatrix}' \begin{bmatrix} c_{ss} & c_{sf} \\ 0 & c_{ff} \end{bmatrix} + \\ \begin{bmatrix} a_{ss} & a_{sf} \\ a_{fs} & a_{ff} \end{bmatrix}' \begin{bmatrix} \varepsilon_{s,t-1}^2 & \varepsilon_{s,t-1}\varepsilon_{f,t-1} \\ \varepsilon_{f,t-1}\varepsilon_{s,t-1} & \varepsilon_{f,t-1}^2 \end{bmatrix} \begin{bmatrix} a_{ss} & a_{sf} \\ a_{fs} & a_{ff} \end{bmatrix} + \\ + \begin{bmatrix} b_{ss} & b_{sf} \\ b_{fs} & b_{ff} \end{bmatrix}' \begin{bmatrix} \sigma_{ss,t-1}^2 & \sigma_{sf,t-1}^2 \\ \sigma_{fs,t-1}^2 & \sigma_{ff,t-1}^2 \end{bmatrix} \begin{bmatrix} b_{ss} & b_{sf} \\ b_{fs} & b_{ff} \end{bmatrix}'$$

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$$\varepsilon_t | \phi_{t-1} \sim BN(0, H_t)$$
 with $\varepsilon_t = \begin{bmatrix} \varepsilon_{st} \\ \varepsilon_{ft} \end{bmatrix}$ and $H_t = \begin{bmatrix} \sigma_{ss,t}^2 & \sigma_{sf,t}^2 \\ \sigma_{fs,t}^2 & \sigma_{ff,t}^2 \end{bmatrix}$

• $\varepsilon_t | \phi_{t-1} \sim t(0, H_t, v)$; where v is the degrees of freedom parameter of a conditional bivariate t-student distribution.

• 2) Diagonal BEKK

$$\begin{aligned} \sigma_{ss,t}^2 &= c_{ss} + a_{ss}\varepsilon_{s,t-1}^2 + b_{ss}\sigma_{ss,t-1}^2 \\ \sigma_{sf,t}^2 &= c_{sf} + a_{sf}\varepsilon_{s,t-1}\varepsilon_{f,t-1} + b_{sf}\sigma_{sf,t-1}^2 \\ \sigma_{ff,t}^2 &= c_{ff} + a_{ff}\varepsilon_{f,t-1}^2 + b_{ff}\sigma_{ff,t-1}^2 \end{aligned}$$

- $\varepsilon_{t} | \phi_{t-1} \sim BN(0, H_{t})$
- $\varepsilon_t | \phi_{t-1} \sim t(0, H_t, v)$; where v is the degrees of freedom parameter of a conditional bivariate t-student distribution.

Methodology: Time varying hedge ratio

- 3) Bollerslev's CCC
- In the Bollerslev's (1990) model, covariances between *i* and *j* are allowed to vary only through the product of standard deviations with a correlation coefficient which is constant through time (constant correlation model or CCC). The dynamic of standard deviations is governed by the GARCH(1,1) variances' dynamic or any univariate GARCH model. Keeping the covariance matrix $\Sigma_t = [\sigma_{ij,t}]$, we have

$$\sigma_{ii,t} = \omega_{ii} + \beta_{ii}\sigma_{ii,t-1} + \alpha_{ii}\eta_{i,t}^2$$

and

$$\sigma_{ij,t} = \rho_{ij} \sqrt{\sigma_{ii,t} \sigma_{jj,t}}$$

Methodology: Time varying hedge ratio

• 4) DCC model of Engle

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- Correlations between returns may not be constant trough time
- The general form of the dynamic conditional correlation (DCC) model introduced by Engle (2002) is defined by

$$\Sigma_t = D_t R_t D_t$$

where $R_t = Q_t^{*-1} Q_t Q_t^{*-1}$
and $Q_t = \left(1 - \sum_{p=1}^P \alpha_p - \sum_{q=1}^Q \beta_q\right) \overline{Q} + \sum_{p=1}^P \alpha_p \left(\eta_{t-p} \eta_{t-p'}'\right) + \sum_{q=1}^Q \beta_q Q_{t-q}$

where D_t is a $n \times n$ diagonal matrix of time varying standard deviations, D_t is a $n \times n$ time varying correlation matrix, \overline{Q} is the unconditional covariance matrix using standardized residuals from the univariate estimates, and Q_t^* is a diagonal matrix of the square root of the diagonal elements of Q_t . Time varying correlation matrix defined as $R_t = \left[\rho_{ij,t}\right]$ with $\left[\rho_{ij,t}\right] = \frac{q_{ij,t}}{\sqrt{q_{ii}q_{ij}}}$. DCC differs from CCC mainly in that it allows the correlation matrix to be changed over time.

Data and Descriptive Statistics

- June 24, 2005 to October 9, 2009 for futures ECX and Bluenext spot
- Oaily returns
- One future contract to hedge the spot price variation: Future Contract Maturing in December 2009 (FutDec09) to cover the all period

ECX Series	mean	variance	skewness	kurtosis					
Spot CO2	0.044	4.045	0.671	45.072					
FutDec05	0.132	2.831	-1.811	12.494					
FutDec06	-0.223	4.864	-0.292	44.226					
FutDec07	-0.918	7.423	-0.821	18.152					
FutDec08	0.110	2.944	-1.558	10.310					
FutDec09	-0.009	3.353	-1.718	20.844					
FutDec10	-0.002	3.322	-1.660	20.104					
FutDec11	0.005	3.335	-1.600	18.576					
FutDec12	0.011	3.404	-1.564	16.965					

Table 1: Descriptive Statistics

Estimated spot and futures volatility for each multivariate model

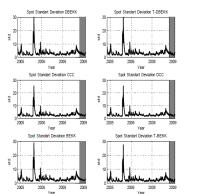
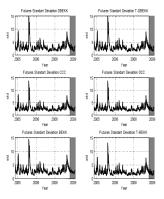


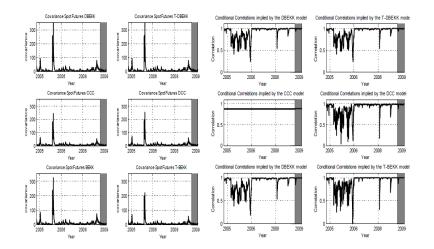
Figure 1: Conditional volatility for the spot CO_2 allowances in the ECX market

Figure 2: Conditional volatility for the Futures December 2009 CO_2 allowances in the ECX market



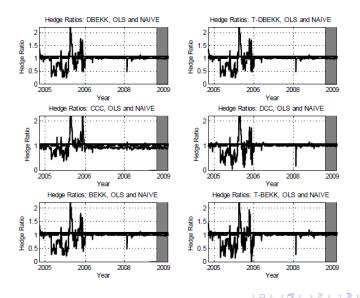
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Estimated Covariance and Conditional Correlation



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Conditional Hedge Ratio



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	In the Sample	Out of Sample
Spot variance (no hedging) $(h = 0)$	13.61	4.53
Hedging	Risk reduction	Risk reduction
Naive $(h = 1)$	74.32	98.62
OLS $\left(h = \frac{\sigma_{FS}}{\sigma_F^2}\right)$	74.04	99.04
Diag-BEKK $\left(h_t = \frac{\sigma_{FS,t}}{\sigma_{F,t}^2}\right)$	85.53	99.04
T-Diag-BEKK $\left(h_t = \frac{\sigma_{FS,t}}{\sigma_{F,t}^2}\right)$	86.63+	99.06
$\text{CCC}\left(h_t = \frac{\sigma_{FS,t}}{\sigma_{F,t}^2}\right)$	83.72	97.72
$\begin{array}{l} \text{CCC} \left(h_t = \frac{\sigma_{FS,t}}{\sigma_{F,t}^2} \right) \\ \text{DCC} \left(h_t = \frac{\sigma_{FS,t}}{\sigma_{F,t}^2} \right) \end{array}$	85.64	99.01
BEKK $\left(h_t = \frac{\sigma_{FS,t}}{\sigma_{Tt}^2}\right)$	85.34	99.00
T-BEKK $\left(h_t = \frac{\sigma_{FS,t}}{\sigma_{F,t}^2}\right)$	85.03	99.07+

Utility Gains for alternative risk aversion levels and different models

 $\eta = 1$; Risk Averse

	In Sample			Out Sample				
	Variance	Return	Exp Util ^a	Gain ^b	Variance	Return	Exp Util ^a	Gain ^b
$\eta = 1$								
Unhedge	13.61	-0.01	-13.63	-	4.53	-0.13	-4.67	_
Naive	3.50	0.00	-3.50	10.13	0.06	0.01	0.00	4.67
OLS	3.53	0.00	-3.53	10.09	0.04	0.00	-0.04	4.62
Diag-BEKK	1.97	0.01	-1.96	11.66	0.04	-0.00	-0.04	4.62
T-Diag-BEKK	1.82	0.00	-1.82	11.81	0.04	-0.00	-0.04	4.63
CCC	2.22	-0.01	-2.23	11.40	0.10	-0.02	-0.12	4.55
DCC	1.96	-0.00	-1.96	11.67	0.04	-0.00	-0.05	4.62
BEKK	2.00	0.01	-1.99	11.64	0.05	-0.00	-0.05	4.62
T-BEKK	2.04	0.00	-2.03	11.59	0.04	-0.00	-0.04	4.63

Utility Gains for alternative risk aversion levels and different models

 $\eta = 2$; Risk Neutral

		In Sample			Out Sample			
	Variance	Return	Exp Util ^a	Gain ^b	Variance	Return	Exp Util ^a	Gain
$\eta = 2$								
Unhedge	13.61	-0.01	-27.24	-	4.53	-0.13	-9.20	-
Naive	3.50	0.00	-6.99	20.25	0.06	0.01	-0.01	9.19
OLS	3.53	0.00	-7.07	20.17	0.04	0.00	-0.09	9.12
Diag-BEKK	1.97	0.01	-3.93	23.31	0.04	-0.00	-0.09	9.12
T-Diag-BEKK	1.82	0.00	-3.64	23.60	0.04	-0.00	-0.09	9.12
CCC	2.22	-0.01	-4.44	22.80	0.10	-0.02	-0.22	8.98
DCC	1.96	-0.00	-3.91	23.33	0.04	-0.00	-0.09	9.11
BEKK	2.00	0.01	-3.98	23.26	0.05	-0.00	-0.09	9.11
T-BEKK	2.04	0.00	-4.07	23.17	0.04	-0.00	-0.08	9.12

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Utility Gains for alternative risk aversion levels and different models

 $\eta =$ 4; Risk Lover

	In Sample			Out Sample				
	Variance	Return	Exp Util ^a	Gain ^b	Variance	Return	Exp Util ^a	Gain ^b
$\eta = 4$								
Unhedge	13.61	-0.01	-54.47	-	4.53	-0.13	-18.27	-
Naive	3.50	0.00	-13.98	40.49	0.06	0.01	-0.03	18.25
OLS	3.53	0.00	-14.14	40.34	0.04	0.00	-0.17	18.10
Diag-BEKK	1.97	0.01	-7.87	46.60	0.04	-0.00	-0.17	18.10
T-Diag-BEKK	1.82	0.00	-7.28	47.19	0.04	-0.00	-0.17	18.10
CCC	2.22	-0.01	-8.88	45.59	0.10	-0.02	-0.43	17.84
DCC	1.96	-0.00	-7.83	46.64	0.04	-0.00	-0.18	18.09
BEKK	2.00	0.01	-7.98	46.49	0.05	-0.00	-0.18	18.09
T-BEKK	2.04	0.00	-8.15	46.32	0.04	-0.00	-0.17	18.10

^a Expected Utility: $E\left[U\left(r_{p,t}\right)|\psi_{t-1}\right] = E\left[r_{p,t}|\psi_{t-1}\right] - \lambda Var\left(r_{p,t}|\psi_{t-1}\right)$

^b Utility Gain of Hedging Models over Unhedged Position

- Empirically estimate optimal hedge ratios in the EU ETS CO₂ allowances markets
- Results indicate that taking into account transaction costs of rebalancing daily the hedged portfolio in dynamic MGARCH models will imply that their better statistical performance in the EU ETS market becomes seriously questioned.
- Taking into account the data leptokurtosis through the error distribution assumption indicates superior gains, measured by variance reduction, obtained from the multivariate model BEKK (Diagonal), for both in sample and out of sample results (BEKK). Moreover, utility gains increase with the investor's preference over risk.

- Overall, there seems to be some gains from including heteroscedasticity and time-varying variances in hedge ratios calculations, although it is not completely guaranteed that improving statistical price modelling provides better performance.
- Correlation results are important for EU ETS allowances price risk management, as they show that December Futures will provide a good risk reduction for hedgers participating in EU ETS markets.
- As the market evolves and more data becomes available, it is expected more useful results obtained through dynamic models or even others given that empirical research is evolving constantly.

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Thanks for your attention! Questions?